

# Chapter #4 Elementary Statistics

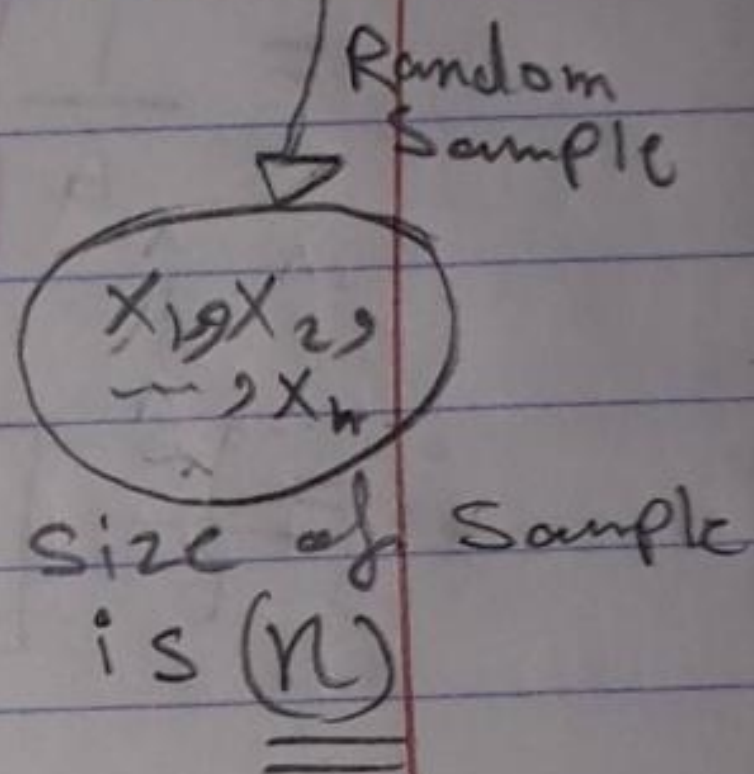
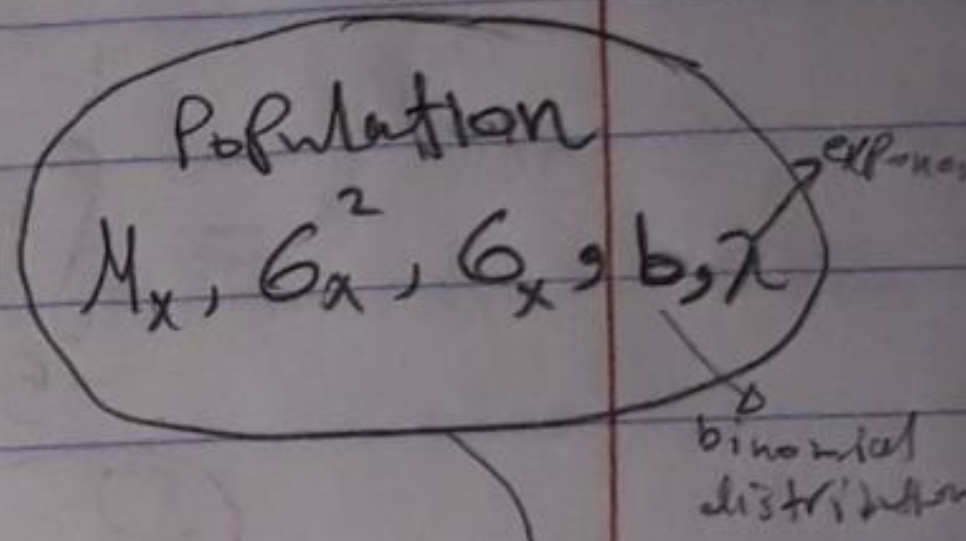
Basic Definitions about the random sample:-

1 Sample mean:  $\hat{\mu}_x = \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$

2 Sample Variance:  $S_x$  or  $\hat{\sigma}_x^2$

$S_x = \hat{\sigma}_x^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu_x)^2$   $\mu_x$  is known (true mean is known).

OR  $\hat{S}_x = \hat{\sigma}_x^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \hat{\mu}_x)^2$  True Mean is unknown so we use  $\hat{\mu}_x$



3 Sample Standard deviation  $S_x, \hat{\sigma}_x$

$S_x = \hat{\sigma}_x = \sqrt{S_x^2} = \sqrt{\hat{\sigma}_x^2}$

∴ Some Relation: give 2 random variables

Sample Covariance =  $\hat{\mu}_{x,y} = C_{x,y}$

$C_{x,y} = \frac{1}{n-1} \sum_{i=1}^n (X_i - \hat{\mu}_x)(Y_i - \hat{\mu}_y)$

Sample Correlation Coefficient =  $r_{xy}$

$r_{xy} = \frac{C_{xy}}{S_x S_y}$

برضوی ما اینتا تباہونہ ال correlation coefficient بتسایر 3

پس ہوں اختلاف الرموز

من صيغة ال Variance  $(\hat{S}_x^2)$  <sup>التالية</sup> بنا نشق صيغة جديدة =

$$S_x^2 = \hat{\sigma}_x^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\mu}_x)^2$$

$$= \frac{1}{n-1} \sum_{i=1}^n [x_i^2 - 2x_i \hat{\mu}_x + \hat{\mu}_x^2]$$

فتبينا التربيع ←

$$= \frac{1}{n-1} \left[ \sum_{i=1}^n x_i^2 - 2\hat{\mu}_x \sum_{i=1}^n x_i + \hat{\mu}_x^2 \sum_{i=1}^n 1 \right]$$

$$= \frac{1}{n-1} \left[ \sum_{i=1}^n x_i^2 - 2\hat{\mu}_x n \hat{\mu}_x + \hat{\mu}_x^2 n \right]$$

$$= \frac{1}{n-1} \left[ \sum_{i=1}^n x_i^2 - 2n \hat{\mu}_x^2 + n \hat{\mu}_x^2 \right]$$

$$= \frac{1}{n-1} \left[ \sum_{i=1}^n x_i^2 - n \hat{\mu}_x^2 \right]$$

$$= \frac{1}{n-1} \left[ \sum_{i=1}^n x_i^2 - n \left( \frac{1}{n} \sum_{i=1}^n x_i \right)^2 \right]$$

∴  $S_x^2 = \frac{1}{n(n-1)} \left[ n \sum_{i=1}^n x_i^2 - \left( \sum_{i=1}^n x_i \right)^2 \right]$

بستعملنا إذاً صيغة  $\hat{\mu}_x$  و  $\mu_x$

أما هذا فمن صيغة ال Covariance  $\hat{S}_{xy}$  <sup>التالية</sup> بنا نشق صيغة جديدة =

$$C_{xy} = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\mu}_x)(y_i - \hat{\mu}_y)$$

$$= \frac{1}{n-1} \sum_{i=1}^n [x_i y_i - x_i \hat{\mu}_y - y_i \hat{\mu}_x + \hat{\mu}_x \hat{\mu}_y]$$

$$= \frac{1}{n-1} \left[ \sum_{i=1}^n x_i y_i - \hat{\mu}_y \sum_{i=1}^n x_i - \hat{\mu}_x \sum_{i=1}^n y_i + \hat{\mu}_x \hat{\mu}_y \sum_{i=1}^n 1 \right]$$

$$= \frac{1}{n-1} \left[ \sum_{i=1}^n x_i y_i - \hat{\mu}_y n \hat{\mu}_x - n \hat{\mu}_y \hat{\mu}_x + n \hat{\mu}_x \hat{\mu}_y \right]$$

$$= \frac{1}{n-1} \left[ \sum_{i=1}^n x_i y_i - \frac{1}{n} \left( \sum_{i=1}^n x_i \right) \left( \sum_{i=1}^n y_i \right) \right]$$

$$C_{xy} = \frac{1}{n(n-1)} \left[ n \sum_{i=1}^n x_i y_i - \left( \sum_{i=1}^n x_i \right) \left( \sum_{i=1}^n y_i \right) \right]$$

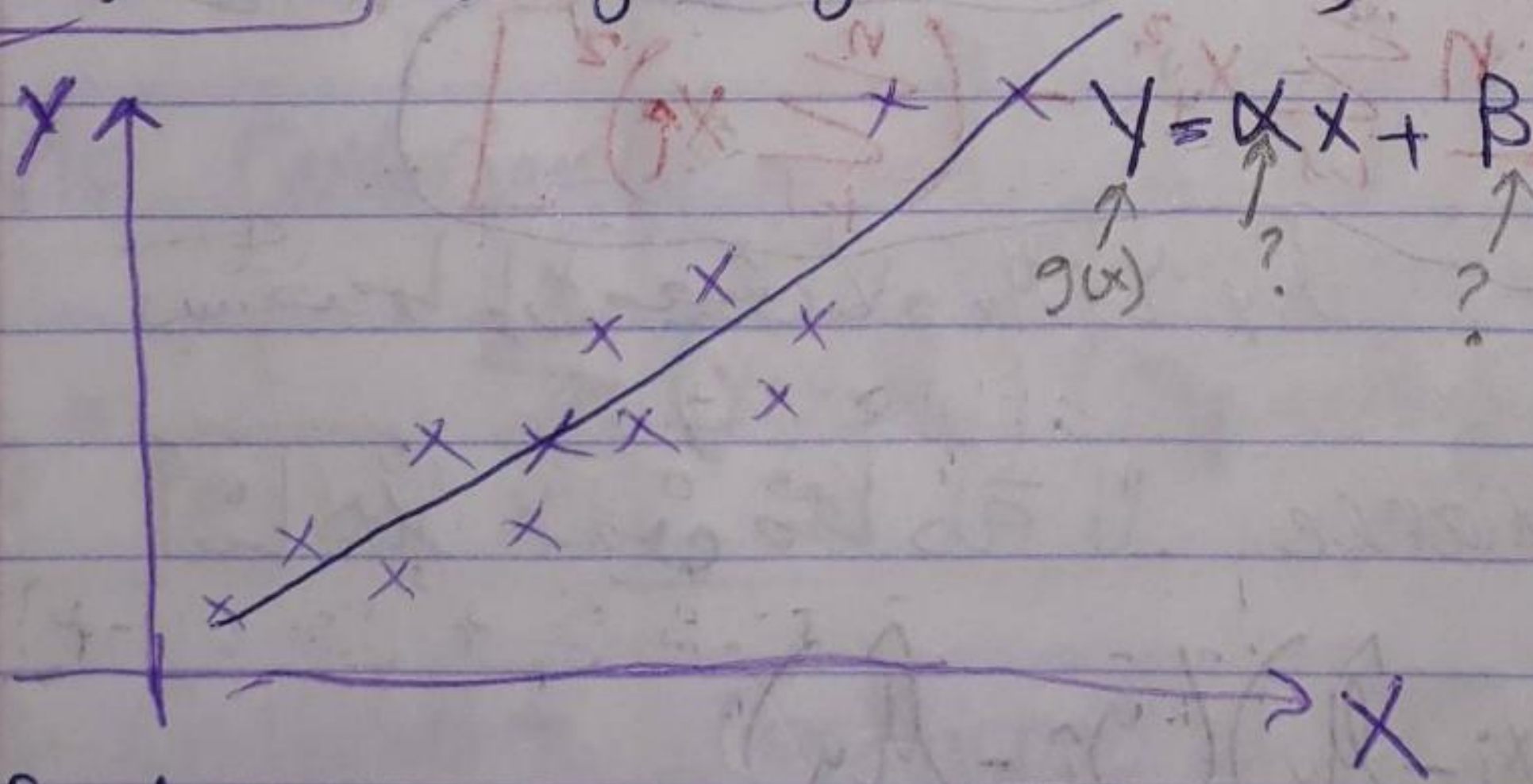
نلاحظ اننا اذا حوطينا  $y$  بـ  $x$  ، فنتبع نفس قانون الـ  $S_x^2$  الى استنتاجه هذا .

Regression Techniques :- two random variables

$x_0$	$x_1$	$x_2$	$x_3$						$x_n$
$y_0$	$y_1$	$y_2$	$y_3$						$y_n$ ← Practical measurement
$y = g(x)$	$g(x)$	$g(x)$	$g(x)$						$g(x_n)$ ← theoretical

Random Sample  
 $\langle x_1, y_1 \rangle, \langle x_2, y_2 \rangle, \dots, \langle x_n, y_n \rangle$

draw the plot out of these points  
 So we try to find the best fit line passing through the points, and Equation.



to find the equation that describes the relationship between  $x$  and  $y$ ; a  $g(x)$  function  $= y$  needs to be defined here.



$$y = \alpha x^2 + \beta x^0 \quad \text{المعادلة الكعبيّة}$$

$$\beta n + \alpha \sum_{i=1}^n x_i = \sum_{i=1}^n y_i \quad \text{--- ①}$$

$$\beta \sum_{i=1}^n x_i + \alpha \sum_{i=1}^n x_i^2 = \sum_{i=1}^n x_i y_i \quad \text{--- ②}$$

التي حوّثت معادلة أقل خطية أول  
 التي حوّثت معادلة أقل خطية بـ  $\beta$  بالمعادلة الأولى  
 التي حوّثت معادلة أقل خطية بـ  $\alpha$  بالمعادلة ②

$$\begin{bmatrix} n & \sum_{i=1}^n x_i \\ \sum_{i=1}^n x_i & \sum_{i=1}^n x_i^2 \end{bmatrix} \begin{bmatrix} \beta \\ \alpha \end{bmatrix} = \begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n x_i y_i \end{bmatrix}$$

نتائج معادلة ①  
 نتائج معادلة ②

$$\beta = \frac{\begin{vmatrix} \sum y_i & \sum x_i \\ \sum x_i y_i & \sum x_i^2 \end{vmatrix}}{\begin{vmatrix} n & \sum x_i \\ \sum x_i & \sum x_i^2 \end{vmatrix}} = \frac{\sum y_i \sum x_i^2 - \sum x_i \sum x_i y_i}{n \sum x_i^2 - (\sum x_i)^2}$$

تكملة

$$\alpha = \frac{\begin{vmatrix} n & \sum y_i \\ \sum x_i & \sum x_i y_i \end{vmatrix}}{\begin{vmatrix} n & \sum x_i \\ \sum x_i & \sum x_i^2 \end{vmatrix}} = \frac{n \sum x_i y_i - \sum x_i \sum y_i}{n \sum x_i^2 - (\sum x_i)^2}$$

تلاحظ ان  $S_x^2 = \frac{1}{n-1} (\sum x_i^2 - \frac{(\sum x_i)^2}{n})$   $\rightarrow$  أربع لمعادلة  $C_{xy}$   $\rightarrow$   $= \frac{n(n-1) C_{xy}}{n(n-1) S_x^2}$

$$\alpha = \frac{-C_{xy}}{S_x^2}$$

$$\begin{aligned}
 y &= \alpha x + \beta \\
 \hat{M}_y &= \alpha \hat{M}_x + \beta \\
 \therefore \beta &= \hat{M}_y - \alpha \hat{M}_x
 \end{aligned}$$

بؤرب اعمى كمان

So we can rewrite the equation  $y = \alpha x + \beta$  as follows:

$$Y = \frac{C_{xy}}{S_x^2} X + \hat{M}_y - \frac{C_{xy}}{S_x^2} \hat{M}_x$$

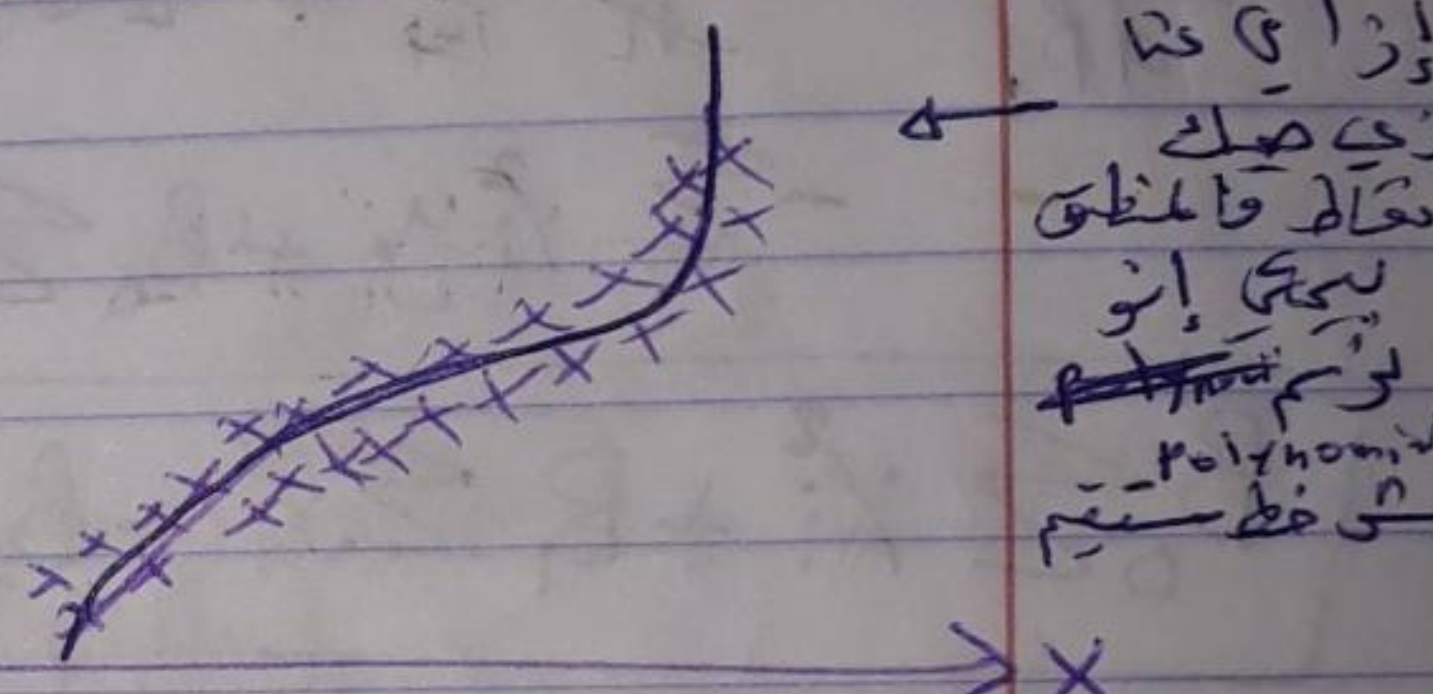
$$Y - \hat{M}_y = \frac{C_{xy}}{S_x^2} (X - \hat{M}_x) \xrightarrow[\text{المعادلة على } S_y]{\text{نضرب طرفيها}} \frac{Y - \hat{M}_y}{S_y} = \frac{C_{xy}}{S_x S_y} \left[ \frac{X - \hat{M}_x}{S_x} \right]$$

$$\therefore \left[ \frac{Y - \hat{M}_y}{S_y} = r_{xy} \left( \frac{X - \hat{M}_x}{S_x} \right) \right]$$

## Polynomial Regression :-

$$Y = \underline{\underline{B_0}} X^0 + \underline{\underline{B_1}} X^1 + \underline{\underline{B_2}} X^2$$

$x_i$	$x_1$	$x_2$	...	$x_n$
$y_i$	$y_1$	$y_2$	...	$y_n$
$g(x)$	$g(x_1)$	$g(x_2)$	...	$g(x_n)$



$$E = \frac{1}{n} \sum_{i=1}^n (y_i - g(x))^2$$

$$E = \frac{1}{n} \sum_{i=1}^n (y_i - [B_0 + B_1 x_i + B_2 x_i^2])^2$$

$$E = \frac{1}{n} \sum_{i=1}^n [y_i - B_0 - B_1 x_i - B_2 x_i^2]^2$$

To find  $B_0, B_1, & B_2 \rightarrow$  DERIVE!

$$\frac{dE}{dB_0} = \frac{1}{n} \sum_{i=1}^n 2 [y_i - B_0 - B_1 x_i - B_2 x_i^2] (-1) = 0$$

$$\Rightarrow -\sum_{i=1}^n y_i + n B_0 + B_1 \sum_{i=1}^n x_i + B_2 \sum_{i=1}^n x_i^2 = 0$$

QED

$$nB_0 + B_1 \sum x_i + B_2 \sum x_i^2 = \sum y_i \quad \text{--- (1)}$$

Derive in respect to  $B_1$  :-

$$\frac{dE}{dB_1} = \frac{1}{n} \sum_{i=1}^n 2 [y_i - B_0 - B_1 x_i - B_2 x_i^2] (-x_i) = 0 \quad \times n$$

$$\therefore -\sum x_i y_i + B_0 \sum x_i + B_1 \sum x_i^2 + B_2 \sum x_i^3 = 0$$

$$\Rightarrow B_0 \sum x_i + B_1 \sum x_i^2 + B_2 \sum x_i^3 = \sum x_i y_i \quad \text{--- (2)}$$

Derive in respect to  $B_2$  :-

$$\frac{dE}{dB_2} = \frac{1}{n} \sum_{i=1}^n 2 [y_i - B_0 - B_1 x_i - B_2 x_i^2] (-x_i^2) = 0$$

$$-\sum x_i^2 y_i + B_0 \sum x_i^2 + B_1 \sum x_i^3 + B_2 \sum x_i^4 = 0$$

$$B_0 \sum x_i^2 + B_1 \sum x_i^3 + B_2 \sum x_i^4 = \sum x_i^2 y_i \quad \text{--- (3)}$$

Using linear algebra :-

$$\begin{bmatrix} n & \sum x_i & \sum x_i^2 \\ \sum x_i & \sum x_i^2 & \sum x_i^3 \\ \sum x_i^2 & \sum x_i^3 & \sum x_i^4 \end{bmatrix} \begin{bmatrix} B_0 \\ B_1 \\ B_2 \end{bmatrix} = \begin{bmatrix} \sum y_i \\ \sum x_i y_i \\ \sum x_i^2 y_i \end{bmatrix}$$

ويجوز الحل باستخدام المصفوفات زي ما حاله  $\alpha$  و  $\beta$  بالسؤال القبل.

# Fitting an Exponential by the Method of Least Squares

If we have the equation:  $y = a e^{bx}$ , to solve it easily:-

We use the Linearization Method:-  
 :- linear function,  $y = \beta_0 + \beta_1 x$

Take the  $(\ln)$  for both sides:-

$$\ln[y] = \ln[a e^{bx}]$$

$$= \ln[a] + \ln[e^{bx}]$$

$$\ln[y] = \ln[a] + bx$$

$$y_{\text{new}} = \beta_0 + \beta_1 x$$

كبرنا المتغير  
replaced

$$\Rightarrow \boxed{\beta_0 = \ln a} \quad \boxed{\beta_1 = b} \quad \boxed{y_{\text{new}} = \ln[y]}$$

و بعد ما نحلها نصل الى  $y = a e^{bx}$

**Ex** If  $y = \frac{L}{1 + e^{a+bx}}$

$$y(1 + e^{a+bx}) = L \Rightarrow y + y e^{a+bx} = L$$

$$e^{a+bx} = \frac{L-y}{y}$$

take  $(\ln)$  for both sides

$$a + bx = \ln\left(\frac{L-y}{y}\right)$$

$$\boxed{y_{\text{new}} = \alpha + \beta x} \rightarrow \text{linear} \quad , \quad y_{\text{new}} = \ln\left(\frac{L-y}{y}\right) \quad , \quad \alpha = a, \beta = b$$



# Central Limit Theorem

Sample mean probability  
population  
 $P(\hat{\mu}_x \leq 11) = ??$

Note:-  $Y = C_1 X_1 + C_2 X_2 + C_3 X_3$

$$E\{Y\} = C_1 \mu_{X_1} + C_2 \mu_{X_2} + C_3 \mu_{X_3}$$

$$\text{Var}\{Y\} = C_1^2 \sigma_{X_1}^2 + C_2^2 \sigma_{X_2}^2 + C_3^2 \sigma_{X_3}^2$$

$$+ 2C_1 C_2 \sigma_{X_1} \sigma_{X_2} \rho_{X_1, X_2}$$

$\rho_{X_1, X_2}$

$$+ 2C_1 C_3 \sigma_{X_1} \sigma_{X_3} \rho_{X_1, X_3}$$

$\rho_{X_1, X_3}$

correlation coefficient

$$+ 2C_2 C_3 \sigma_{X_2} \sigma_{X_3} \rho_{X_2, X_3}$$

$\rho_{X_2, X_3}$

\* if  $X_1, X_2$  and  $X_3$  are S.I, then

$$\rho_{X_1, X_2} = \rho_{X_1, X_3} = \rho_{X_2, X_3} = 0$$

independent

**Ex** Let  $X_1$  and  $X_2$  be two Gaussian random variables such that :-  $\mu_1 = 0, \sigma_1^2 = 4, \mu_2 = 10, \sigma_2^2 = 9, \rho_{12} = 0.25$ . Define  $Y = 2X_1 + 3X_2$ .

**Q** Find the mean and variance of  $Y$ .

$$\begin{aligned} \mu_Y &= 2\mu_{X_1} + 3\mu_{X_2} \\ &= 2(0) + 3(10) \\ &= 30 \end{aligned}$$

$$\begin{aligned} \sigma_Y^2 &= (2)^2 \sigma_{X_1}^2 + (3)^2 \sigma_{X_2}^2 + 2(2)(3) \sigma_{X_1} \sigma_{X_2} \rho_{X_1, X_2} \\ &= 4(4) + 9(9) + 2\sqrt{4}\sqrt{9}(0.25) \\ &= 115 \end{aligned}$$

b) Find  $P(Y < 35)$ .  $Y$  is Gaussian ✓  
 $\mu_Y = 30$ ,  $\sigma_Y^2 = 115$

$$\therefore \phi\left(\frac{35-30}{\sqrt{115}}\right) = \phi(0.466) = 0.6794$$

Ex] Let  $X_1$  and  $X_2$  be two independent Gaussian Random Variables, such that:  $\mu_1 = 0$ ,  $\sigma_1^2 = 4$ ,  $\mu_2 = 10$ ,  $\sigma_2^2 = 9$ . Define  $Y = 2X_1 + 3X_2$ .

a) Find the mean and variance of  $Y$ .

$$\begin{aligned} \mu_Y &= 2\mu_{X_1} + 3\mu_{X_2} \\ &= 2(0) + 3(10) \\ &= 30 \end{aligned}$$

$$\begin{aligned} \sigma_Y^2 &= 4\sigma_{X_1}^2 + 9\sigma_{X_2}^2 + 2 \cancel{2 \cdot 3 \cdot 0} \\ &= 4(4) + 9(9) \\ &= 97 \end{aligned}$$

b)  $P(Y < 35)$  → Gaussian ✓,  $\mu_Y = 30$ ,  $\sigma_Y^2 = 97$   
 $P(Y < 35) = \phi\left(\frac{35-30}{\sqrt{97}}\right) = \phi(0.5077) = 0.6942$

Ex] Soft drink cans are filled by an automated filling machine. The mean fill volume is 330 ml, and the standard deviation is 1.5 ml. Assume that the fill volume of the cans are independent Gaussian Random Variables. What is the probability that the average volume of 10 cans selected at random from this process is less than 328 ml.  
 $P(\bar{M}_x \leq 328) = ??$

$\mu_x = 330 \text{ ml}$  Population  
 $\sigma_x = 1.5$  cans

As  $P(\hat{\mu}_x \leq 328) = ??$   
 we deal with  $\mu_x$  as if it is  $Y$

$X_1, X_2, \dots, X_{10}$   $n=10$

$$\therefore \hat{\mu}_x = Y = \frac{1}{n} \sum X_i$$

$$= \frac{1}{n} X_1 + \frac{1}{n} X_2 + \frac{1}{n} X_3 + \dots + \frac{1}{n} X_{10}$$

$$E\{Y\} = E\{\hat{\mu}_x\} = E\left\{ \frac{1}{n} X_1 + \frac{1}{n} X_2 + \frac{1}{n} X_3 + \dots + \frac{1}{n} X_{10} \right\}$$

$$\therefore E\{\hat{\mu}_x\} = \frac{1}{n} E\{X_1\} + \frac{1}{n} E\{X_2\} + \frac{1}{n} E\{X_3\} + \dots + \frac{1}{n} E\{X_{10}\}$$

$$= \frac{1}{n} \mu_x + \frac{1}{n} \mu_x + \frac{1}{n} \mu_x + \dots + \frac{1}{n} \mu_x$$

$$= \frac{1}{n} (n) (\mu_x)$$

$$= \mu_x$$

$$= 330$$

عندما ال  $X$   $\mu$  ما توجد من ال  
 Population  $\mu$  فال  $E$  يتقارب  $X$  تقريبا  
 ال  $E$  الموجود بال Population

\* يتقارب ال Limit Theorem :- إذا باقر Sample من Population ال

ال  $E$  ال mean الحقيقي  $E\{\hat{\mu}_x\}$  + Sample  $n$  expected value Mean

$$E\{\hat{\mu}_x\} = \mu_x$$

$$\text{Var}\{\hat{\mu}_x\} = \left(\frac{1}{n}\right)^2 \sigma_{x_1}^2 + \left(\frac{1}{n}\right)^2 \sigma_{x_2}^2 + \dots + \left(\frac{1}{n}\right)^2 \sigma_{x_{10}}^2$$

$$= \frac{1}{n^2} (n) \sigma_x^2 = \frac{\sigma_x^2}{n} = \frac{(1.5)^2}{10} = 0.225$$

Gaussian  $\sigma$  ال  $\mu_x = 330$   $\sigma_x^2 = 0.225$

$$P(\hat{\mu}_x < 328) = \Phi\left(\frac{328 - 330}{\sqrt{0.225}}\right) = \Phi\left(\frac{-2}{0.4743}\right) = \Phi(-4.2163)$$

ربط ال الجواب مع الجرد

**Ex 5-10** An Electric company manufactures resistors that have a mean resistance of  $100 \Omega$  and a standard deviation of  $10 \Omega$ . Find the probability that a random sample of  $n=25$  resistors will have an average resistance less than  $95 \Omega$ .

$\mu_x = 100 \Omega$  ,  $\sigma_x = 10$  ,  $n = 25$

Random Sample  
 از آنجمله رندوم سامله  
 که از آن نمونه‌ها  
 independent.

$P(\hat{\mu}_x < 95) = ??$

Using the Central theorem :-

$\hat{\mu}_x = \mu_x = 100$

$\hat{\sigma}_x^2 = \frac{\sigma_x^2}{n} = \frac{10^2}{25} = 4$

$\therefore P(\hat{\mu}_x < 95) = \Phi\left(\frac{95 - 100}{\sqrt{4}}\right) = \Phi\left(-\frac{5}{2}\right) = 0.00621$

suppose it is a Gaussian  
 $\mu_x = 100$   
 $\sigma_x^2 = 4$

**Ex** The lifetime of a special type of battery is a R.V with mean 40 hours and standard deviation 20 hours. A battery is used until it fails, then it is immediately replaced by a new one. Assume we have 25 such batteries, the lifetime of which are independent, approximate the probability that at least 1100 hours of use can be obtained.

$\mu_x = 40$      $\sigma_x = 20$      $n = 25$

Let  $x_1, x_2, x_3, \dots, x_{25}$  be the lifetimes of batteries.  
 Let  $Y = x_1 + x_2 + \dots + x_{25}$  be the overall lifetime of the system.  
 Since  $x_i$  are independent, Using Gaussian :-

$\mu_y = \mu_1 + \mu_2 + \dots + \mu_{25} = 25 \mu_x = 25 \times 40 = 1000$

$\sigma_y^2 = \sigma_{x_1}^2 + \sigma_{x_2}^2 + \dots + \sigma_{x_{25}}^2 = 25 (20)^2 = 10000$

$P(Y \geq 1100) = 1 - P(Y < 1100) = 1 - \Phi\left(\frac{1100 - 1000}{\sqrt{10000}}\right) = 1 - \Phi\left(\frac{100}{100}\right)$   
 $= 1 - \Phi(1) = 0.158655$

## Estimation of Parameters :-

• تقدير المعاملات :- هو فنون في القوائم الكائنية من أجل

Estimator :- is a function of the observable sample data that is used to estimate an unknown population parameter ( $\mu_x, \sigma_x^2, \dots$ ).

We consider two types of estimators :-  
1. Point Estimator      2. Interval Estimator ✓

### Point Estimator

Properties :- 1 An estimator should be close to the true value of the unknown parameters.

Defo: 1 A point estimator ( $\hat{\theta}$ ) is unbiased estimator of ( $\theta$ ) if  $E(\hat{\theta}) = \theta$ .

\* If the estimator is biased, then  $E(\hat{\theta}) - \theta = B$  is called the bias of the estimator ( $\hat{\theta}$ ).

\* Let  $\hat{\theta}_1, \hat{\theta}_2$  be unbiased estimators of ( $\theta$ ), the one with the smallest variance is called the minimum variance unbiased estimator (MVUE).

### In other words

$E\{\hat{\mu}_x\} = \mu_x$   $\leftarrow$   $\hat{\mu}_x$  is an unbiased estimator for the mean  $\mu_x$ .

$E\{\hat{p}\} = p$   $\leftarrow$   $\hat{p}$  is an unbiased estimator for the probability of success  $p$ .

∴ chapter 2, which is  $E\{\hat{M}_x\}$  حسب

$$E\{X\} \equiv \sum_{-\infty}^{\infty} x P(X=x) \rightarrow \text{PMF}$$

$$\equiv \int_{-\infty}^{\infty} x f_x(x) dx \rightarrow \text{PDF}$$

**[Ex]** Let  $X_1$  and  $X_2$  be a random sample of size two from a population with mean  $\mu_x$  and variance  $\sigma_x^2$ . Two estimators for  $\mu_x$  are proposed:-

$M_1 = \frac{X_1 + X_2}{2}$  and  $M_2 = \frac{X_1 + 2X_2}{3}$  which estimator is better and in what sense?

estimator  
المقدر

$$E\{M_1\} = E\left\{\frac{X_1 + X_2}{2}\right\} = E\left\{\frac{X_1}{2}\right\} + E\left\{\frac{X_2}{2}\right\} = \frac{1}{2}\mu_x + \frac{1}{2}\mu_x$$

والمعنى  
يعني

$$= \mu_x \Rightarrow \text{as } E\{\hat{M}_1\} = \mu_x \Rightarrow \hat{M}_1 = \frac{X_1 + X_2}{2} \text{ is unbiased.}$$

For  $\hat{M}_2 \Rightarrow E\{\hat{M}_2\} = E\left\{\frac{X_1 + 2X_2}{3}\right\} = E\left\{\frac{X_1}{3}\right\} + \frac{2}{3}E\{X_2\}$

$$= \frac{1}{3}\mu_x + \frac{2}{3}\mu_x = \mu_x \rightarrow \hat{M}_2 = \frac{X_1 + 2X_2}{3} \text{ is unbiased}$$

Estimation <sup>تقدير</sup> <sub>في اثنين</sub> :-  $M_{x3} = \frac{X_1 + X_2 + 1}{3}$

$$\therefore E\{M_{x3}\} = E\left\{\frac{1}{3}X_1 + \frac{1}{3}X_2 + \frac{1}{3}\right\} = \frac{1}{3}E\{X_1\} + \frac{1}{3}E\{X_2\} + \frac{1}{3}$$

$$= \frac{2}{3}\mu_x + \frac{1}{3} \rightarrow \text{is a biased estimator}$$

∴  $\hat{M}_2$  is also biased <sup>المقدر</sup>

$$B = E\{\hat{\theta}\} - \theta$$

$$= E\{\hat{M}_x\} - \mu_x$$

$$= \frac{2}{3}\mu_x + \frac{1}{3} - \mu_x = \frac{1}{3} - \frac{1}{3}\mu_x$$

$$\therefore B = \frac{1 - \mu_x}{3}$$

بما أن طرق عينة 1 unbiased و 2 biased ، فأكبر ربح أسبقي  
 unbiased ، وعلى ~ أكثر أنو أفضل من بين ال biased ، بأخذ  
 ال estimator التي ال Variance تكون أقل .

$$\text{Var} \{ \hat{\mu}_{x,2} \} = \text{Var} \left\{ \frac{1}{3}x_1 + \frac{2}{3}x_2 \right\}$$

$$= \frac{1}{9} \sigma_x^2 + \frac{4}{9} \sigma_x^2$$

$$\sigma_{\hat{\mu}_{x,2}}^2 = \frac{5}{9} \sigma_x^2$$

central Theo.   
 إذا كان  $x_1, x_2 \sim SI$  ،  
 $y = c_1 x_1 + c_2 x_2$

$$\therefore \sigma_y^2 = c_1^2 \sigma_x^2 + c_2^2 \sigma_x^2$$

$$\text{Var} \{ \hat{\mu}_{x,1} \} = \text{Var} \left\{ \frac{x_1 + x_2}{2} \right\}$$

$$= \frac{1}{4} \sigma_x^2 + \frac{1}{4} \sigma_x^2 = \frac{1}{2} \sigma_x^2$$

$$\text{As } \text{Var} \{ \hat{\mu}_{x,1} \} < \text{Var} \{ \hat{\mu}_{x,2} \}$$

$\therefore \hat{\mu}_{x,1}$  is the best estimator for  $\mu_x$  .

**Example** Check whether the following estimator  
 is biased or unbiased. Then try to modify  
 the estimator to be unbiased if it is found  
 to be biased.

$$\hat{\sigma}_x^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu}_x)^2$$

$$E \{ \hat{\sigma}_x^2 \} = E \left\{ \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu}_x)^2 \right\} = E \left\{ \frac{1}{n} \sum_{i=1}^n (x_i^2 - 2x_i \hat{\mu}_x + \hat{\mu}_x^2) \right\}$$

$$= \frac{1}{n} E \left\{ \sum_{i=1}^n x_i^2 - 2 \hat{\mu}_x \sum_{i=1}^n x_i + n \hat{\mu}_x^2 \right\}$$

$$= \frac{1}{n} E \left\{ \sum_{i=1}^n x_i^2 - 2 \hat{\mu}_x \cdot n \hat{\mu}_x + n \hat{\mu}_x^2 \right\}$$

$$\frac{\sum x_i}{n} = \hat{\mu}_x$$

$$= \frac{1}{n} E \left\{ \sum_{i=1}^n x_i^2 - n \hat{\mu}_x^2 \right\}$$

$$= \frac{1}{n} \left[ \sum E \{ x_i^2 \} - n E \{ \hat{\mu}_x^2 \} \right]$$

\* Note:  $\text{Var} \{ ? \} = E \{ ?^2 \} - (E \{ ? \})^2$

$$\therefore \text{Var} \{ x \} = E \{ x^2 \} - (E \{ x \})^2$$

$$\therefore \sigma_x^2 = E \{ x^2 \} - \mu_x^2$$

$$\therefore E \{ x^2 \} = \sigma_x^2 + \mu_x^2 \quad \text{--- (1)}$$

∴ According to the previous Note :-

$$\text{Var} \{ \hat{\mu}_x \} = E \{ \hat{\mu}_x^2 \} - (E \{ \hat{\mu}_x \})^2$$

According to Central Limit theorem

$$\frac{\sigma_x^2}{n} = E \{ \hat{\mu}_x^2 \} - \mu_x^2$$

$$\therefore E \{ \hat{\mu}_x^2 \} = \frac{\sigma_x^2}{n} + \mu_x^2 \quad \text{--- (2)}$$

$$\therefore = \frac{1}{n} \left( \sum \sigma_{x_i}^2 + n \mu_x^2 \right) - n \left( \frac{\sigma_x^2}{n} + \mu_x^2 \right)$$

$$= \frac{1}{n} \left( n \sigma_x^2 + n \mu_x^2 - \sigma_x^2 - n \mu_x^2 \right)$$

$$\therefore E \{ \hat{\sigma}_x^2 \} = \frac{1}{n} (n \sigma_x^2 - \sigma_x^2) = \frac{n-1}{n} \sigma_x^2 \rightarrow \text{biased}$$

فإنه غير متحيز  $\hat{\sigma}_x^2 = \sigma_x^2$  وهو المطلوب  $\rightarrow$  unbiased  
 فإنها غير متحيز  $\hat{\sigma}_x^2$  مع  $\sigma_x^2$   $\sim$   $\hat{\sigma}_x^2$  unbiased

$$\therefore E \left\{ \frac{n}{n-1} \hat{\sigma}_x^2 \right\} = E \left\{ \frac{n}{n-1} * \frac{1}{n} \sum (x_i - \hat{\mu}_x)^2 \right\} = \frac{n}{n-1} * \frac{n-1}{n} \sigma_x^2$$

$$\therefore \hat{\sigma}_{x_{new}}^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\mu}_x)^2 \Rightarrow E \{ \hat{\sigma}_{x_{new}}^2 \} = \sigma_x^2$$

لذا طبع القابل 20 بين ضرب الـ 19 بالمجموع وقله فشيء ما لا بد  
 بالبيانات  $\hat{\sigma}_x^2$